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DO NONRECIPROCAL PREFERENCE REGIMES INCREASE EXPORTS?

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Abstract

This paper investigates whether and to what extent nonreciprocal preference regimes have increased developing countries’ exports to richer countries. Moreover, it analyzes how they have affected donors’ exports to beneficiary countries. Using recent developments in the econometric analysis of the gravity equation over the period 1990-2008, we find robust evidence that, on the whole, nonreciprocal preference regimes and GSP schemes have had an economically significant effect on exports from developing countries. However, the estimation of catch-all dummies masks heterogeneous results for the individual schemes. Finally, we find that nonreciprocal regimes have also increased exports from donors to beneficiary countries.

Key words: Nonreciprocal preference regimes; GSP; EBA; AGOA; developing countries.
JEL Classification: F14.

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1. Introduction

The increase of exports from developing countries to industrialized nations’ markets has long been considered an essential element to reduce poverty, promote sustainable development and reap the potential benefits of globalization for the developing world. While there has been an intense debate in policy-making circles on how best to accomplish these aims, the prevailing approach has implied that developed countries give support to the integration of developing countries into the world economy through an “special and differential treatment” (in the form of nonreciprocal preferences) for imports from the developing world. The leading instrument for such trade preferences has been the Generalized System of Preferences, but there exist other unilateral (nonreciprocal) preference regimes that are part of this approach.

The Generalized System of Preferences (GSP) is an exception to the GATT principles of reciprocity and nondiscrimination emerged in second half of the 1960s, through which developed countries provide preferential access to their markets to a large number of developing countries and territories. Australia was the first developed country authorized to establish a GSP for developing countries, and since the early 1970s other developed countries followed in Australia footsteps (EU’s countries, US, Canada or New Zealand, among others).

In addition to the standard GSP schemes, the EU and the US have signed other preference regimes with poor countries. On the one hand, the Cotonou Agreement (also known as ACP-EU Partnership Agreement) is the most comprehensive partnership agreement between developing countries from Africa, the Caribbean and the Pacific (ACP) and the European Union (EU). The basic principle of Cotonou Agreement (henceforth ACP-EU) is that, with some exceptions, the ACP countries’ industrial exports have duty- and quota-free access to the EU market. Another preference regime, that forms part of the EU’s GSP scheme, is the Everything But Arms (EBA) arrangement, which provide unilateral trade preferences to the EU market for products from the 49 Least Developed Countries (LDCs). On the other hand, besides the United
States’ GSP program (that started in 1976), the US administration also grants other, more recent, nonreciprocal preference regimes including the Caribbean Basin Initiative (CBI), the Andean Trade Preference Act (ATPA) and the African Growth and Opportunity Act (AGOA).

This paper investigates whether and to what extent unilateral preference regimes (UPRs) have increased developing countries’ exports. Moreover, it also analyzes how they have affected donors’ own export performance to the corresponding beneficiary countries. GSP and other unilateral regimes are nonreciprocal programs. However, developed countries take into account their own commercial interests in the design of the criteria for eligible countries. For instance, in the case of the US scheme of preferences such criteria include ensuring “equitable and reasonable” access in the beneficiaries’ market to US products, protecting intellectual property rights, and preventing the seizure of property belonging to US citizens or businesses. In fact, following with the same example, the statutory goals of the US GSP include not only the development of developing countries by trade (rather than aid) as a more efficient way of promoting economic growth and development, but also the promotion of US exports in developing countries markets and trade liberalization in developing countries (Jones, 2006).

To the best of our knowledge, there are no studies that neither measure and compare the effect on the developing countries exports of all nonreciprocal preference regimes nor investigate the potential impact in the reverse direction. This paper fits within a larger literature that attempts to measure the effect of policies on bilateral trade using gravity equations.¹ In particular, we estimate the effect of UPRs on exports with several estimation techniques including a recently developed econometric approach: the two-stage estimation procedure proposed by Helpman, Melitz and Rubinstein (2008). This technique allows to

¹ The main branch of that literature examines the effect of trade agreements (see, for example, Baier and Bergstrand, 2007; Baier, Bergstrand and Vidal, 2007; Carrère, 2006; Gil, Llorca and Martínez-Serrano, 2008a or Lee Park and Shin, 2008). But the gravity model has also been regularly used to estimate the trade effects of currency unions (Rose, 2000; Glick and Rose, 2002; Micco, Stein and Ordoñez, 2003), exchange rate regimes (Klein and Shambaugh, 2006 or Gil, Llorca and Martínez-Serrano, 2007) GATT/WTO membership (Rose, 2004; Subramanian and Wei, 2007 or Tomz, Goldstein and Rivers, 2007), and even of the physical presence of government officials in the destination markets or the existence of state visits (Rose, 2007; Nitsch, 2007; Gil, Llorca and Martínez-Serrano, 2008b or Volpe-Martincus and Carballo, 2008).
correct for selection bias and to account for exporter heterogeneity. The sample covers 177 countries over the period 1990-2008.

To preview our results, we find robust evidence that, on the whole, nonreciprocal preference regimes and, in particular, GSP schemes have had an economically significant effect on exports. The ACP-EU, EBA as well as GSP schemes of EU, US, Japan, Canada, Norway, Switzerland and Turkey show a positive effect on developing countries exports to the corresponding developed markets. However, we do not find evidence that membership in AGOA has had a positive effect on exports from African countries to the US. The same occurs for membership in the GSP schemes of Australia, New Zealand or Russia. Finally, we find that nonreciprocal preference regimes have also boosted exports from the donors (developed countries) to the beneficiaries (developing countries).

The paper is structured as follows. Section 2 presents the methodology. Section 3 describes the data. Section 4 discusses the estimation results. Finally, section 5 concludes the paper.

2. Methodology

The international trade literature provides two kinds of approaches to analysing the effects of preferential trade agreements (PTAs). The ex-ante approach, which uses computable general equilibrium (CGE) models, and the ex post approach, which measures trade effects by means of regression techniques. One advantage of CGE models is that they can be used to draw direct inferences about consumption, output and welfare. However, one major limitation of these models is that they use restrictive assumptions and very simple characterisations of real-world preferential trade agreements. In contrast, while the econometric studies cannot analyse consumption, output or welfare effects directly, they have three major advantages: implementation simplicity, superior empirical performance and the possibility of examining actual PTAs.

The gravity equation has emerged as the empirical workhorse in international trade for examining the ex-post effects of PTAs on bilateral trade flows. Therefore, to estimate the ex post effects of the special trade preferences given by developed countries to developing countries on international trade, we
rely on the standard gravity model of trade, which relates bilateral trade flows to economic size, distance and other factors that affect trade barriers.\(^2\)

We estimate the following general equation:

\[
\ln X_{ijt} = \beta_0 + \beta_1 \ln Y_i + \beta_2 \ln Y_j + \beta_3 \ln D_{ij} + \beta_4 \text{Cont}_{ij} + \beta_5 \text{Island}_{ij} \\
+ \beta_6 \text{Landl}_{ij} + \beta_7 \text{Lang}_{ij} + \beta_8 \text{Colony}_{ij} + \beta_9 \text{ComCountry}_{ij} \\
+ \beta_{10} \text{Creligion}_{ij} + \beta_{11} \text{CU}_{ij} + \beta_{12} \text{PTAPlur}_{ij} + \beta_{13} \text{PTABil}_{ij} + \beta_{14} \text{UPR}_{ij} + u_{ijt} \tag{1}
\]

where \(i\) and \(j\) denote trading partners, \(t\) is time, and the variables are defined as follows:

- \(X_{ijt}\) are the bilateral export flows from \(i\) to \(j\) in year \(t\),
- \(Y\) denotes Gross Domestic Product,
- \(D\) denotes the distance between \(i\) and \(j\),
- \(\text{Cont}\) is a dummy variable equal to one when \(i\) and \(j\) share a land border,
- \(\text{Island}\) is the number of island nations in the pair (0, 1, or 2),
- \(\text{Landl}\) is the number of landlocked areas in the country-pair (0, 1, or 2),
- \(\text{Lang}\) is a dummy variable which is unity if \(i\) and \(j\) have a common language,
- \(\text{Colony}\) is a binary variable which is unity if \(i\) ever colonized \(j\) or vice versa,
- \(\text{ComCountry}\) is a binary variable which is unity if \(i\) and \(j\) were part of a same country in the past,
- \(\text{Creligion}\) is an index of common religion\(^3\),
- \(\text{CU}\) is a binary variable which is unity if \(i\) and \(j\) use the same currency at time \(t\),
- \(\text{PTAPlur} (\text{PTABil})\) is a binary variable which is unity if \(i\) and \(j\) belong to the same plurilateral (bilateral) preferential trade agreement,
- \(\text{UPR}\) is a binary variable which is unity if \(i\) is a beneficiary of an Unilateral Preference Regime and \(j\) is the corresponding preference-giving country, and
- \(u_{ijt}\) is the standard classical error term.

We estimate the gravity equation (1) in a number of different ways. We begin with conventional ordinary least squares, including a full set of year-specific intercepts and using robust standard errors. Next we run the gravity equation using both country fixed effects (CFE) and country year fixed effects

\(^2\) The initial applications of the gravity equation to international trade lacked theoretical foundation. However, since the end of the 1970’s the situation has changed and nowadays the gravity equation is backed up by sound theory. See, among others, Anderson (1979), Bergstrand (1985 and 1989), Deardoff (1998), Evenett and Keller (2002), Eaton and Kortum (2002) and Anderson and van Wincoop (2003).

\(^3\) The index is defined as: (% Protestants in country \(i\) * % Protestants in country \(j\)) + (% Catholics in country \(i\) * % Catholics in country \(j\)) + (%Muslims in Country \(i\) * % Muslims in country \(j\)).
The strategy of using CFE sufficiently addresses multilateral resistance in a cross section but CYFE are required to comprehensively control for multilateral resistance in panel datasets (Anderson and van Wincoop, 2003 and 2004). We additionally employ an additional and recently developed econometric approach: the two-stage estimation procedure proposed by HMR (2008), which allows us to correct for selection bias and to account for exporter heterogeneity. This procedure is briefly outlined next.

The HMR (2008) estimation procedure consists in two-stages. In the first stage they estimate a probit equation that specifies the probability that country \( i \) exports to \( j \) conditional on the observable variables. In the second stage, predicted components of this equation are used to estimate the gravity equation. This procedure simultaneously corrects for two types of potential biases: a Heckman selection bias and a bias from potential asymmetries in the trade flows between pairs of countries.

More formally, in a first stage they estimate a probit equation of the type:

\[
\Pr(\text{observed variables}) = \Phi(\chi_i, \lambda_j, X_{ij}, Z_{ij}, \varepsilon_{ij})
\]  

where \( T_{ij} \) is an indicator variable equal to 1 when country \( i \) exports to \( j \) and zero when it does not, \( \Phi \) is the cumulative distribution function of the standard normal distribution, \( \chi_i \) and \( \lambda_j \) are exporter and importer fixed effects, \( X_{ij} \) are variables which affect both the probability and the volume of trade, and \( Z_{ij} \) represents variables that are used for the exclusion restriction, that is, those that affect the probability of observing a positive volume of trade but do not impact the volume of trade if this were to be positive.\(^4\) Using the probit regression, they construct two variables that are included as regressors in the second stage estimation. One is the inverse of Mills ratio and the other is an expression that controls for firm size heterogeneity. In particular, the second stage consists in the estimation for a given year of the following non-linear equation for all country-pairs with positive trade flows:

\(^4\) In this set-up, parameter identification requires the existence of a variable that affects the probability of observing a non-zero flow between two countries but not the volume. Alternatively, a variable which affects both decisions in opposite directions would also work.
\[ \ln \text{trade}_{ij} = \beta_0 + \lambda_j + \chi_i - \gamma X_{ij} + \theta \tilde{\pi}_y + \ln \left\{ \exp \left[ \delta (\tilde{z}_y^* + \tilde{\rho}_y^*) - 1 \right] \right\} + \epsilon_{ij} \]  

(3)

where \( \tilde{\pi}_y \) is the inverse Mills ratio and \( \tilde{z}_y^* = \Phi^{-1}(\tilde{\rho}_y^*) \) in which \( \tilde{\rho}_y \) are the estimates from the probit equation.\(^5\)

3. Data

The trade data for the regressand (export flows from country \( i \) to country \( j \)) come from the “Direction of Trade” (DoT) dataset built up by the International Monetary Fund (IMF). The data comprise bilateral merchandise trade between 177 countries and territories (see Appendix) for seven years of the period 1990-2008 at three-year intervals (1990, 1993, 1996, 1999, 2002, 2005 and 2008).\(^6\) The DoT dataset provides FOB exports in US dollars. These series are converted into constant terms using the American GDP deflator taken from the Bureau of Economic Analysis (US Department of Commerce).

The independent variables come from different sources. GDP data in constant US dollars are taken from the World Development Indicators (World Bank). For location of countries (geographical coordinates), used to calculate Great Circle Distances, and the construction of the dummy variables for physically contiguous neighbours, island and landlocked status, common language, colonial ties, common religion and common country background data are taken from the CIA’s World Factbook. The indicators of preferential trade agreements have been built using data from the World Trade Organization, the Preferential Trade Agreements Database (The Faculty of Law at McGill University) and the web site http://ec.europa.eu/trade/issues/bilateral/index_en.htm. The indicators of currency unions are taken from Reinhart and Rogoff (2002), CIA’s World Factbook and Masson and Pattillo (2005). The sample includes 192 preferential trade agreements (plurilateral and bilateral) and 17 currency unions.\(^7\) Data on

\(^5\) Since equation (3) is non-linear in \( \delta \), following HMR (2008) we estimate it using maximum likelihood.

\(^6\) It is noteworthy that not all the areas considered are countries in the conventional sense of the word. We also include some dependencies, territories and overseas departments in the data.

\(^7\) The expression PTAs in this paper refers also to other agreements involving a higher degree of economic integration. In fact, most economic integration agreements considered in the sample are free trade agreements.
the key variables AGOA and EBA come from the corresponding web pages. The list of beneficiaries of the Cotonou Agreement comes from http://ec.europa.eu/trade/wider-agenda/development/economic-partnerships. The list of countries beneficiaries of the standard GSP schemes are taken from the United Nations Conference on Trade and Development (2008). The list of Caribbean Basin Initiative (CBI) and Andean Trade Preference Act (ATPA) come from the Office of United States Trade Representative.

4. Empirical results

Our benchmark specification to estimate the effect of unilateral preference regimes on developing countries exports is Ordinary Least Squares (with a full set of year-specific intercepts added to correct for common shocks and trends). The results are reported in column 1 of Table 1. The gravity equation works well in two senses. First, the equation fits the data well explaining around two-thirds of the variation of bilateral exports flows. Second, the estimated coefficients are, on the whole, intuitive in sign and size and both economically and statistically significant. The negative effect of a common religion is the exception. Economically larger countries trade more and more distant countries trade less. Landlocked countries trade less, whereas sharing a common border, a common language, a common currency, or sharing membership in a plurilateral or bilateral preferential trade agreements increase trade. The existence of colonial ties encourages trade, as do being islands or part of the same country in the past. With regard to the variable of interest (UPR), we find an estimated coefficient that is positive but statistically insignificant at conventional levels.

Columns 2 and 3 of Table 1 contains regression results adding country-specific fixed effects (CFE) and country year fixed effects (CYFE) to the benchmark equation, respectively. The inclusion of CFE (column 2) controls for the multilateral resistance terms under the assumption that these terms do not vary over time. In almost all cases, the impact goes in the same direction than in column 1. The exceptions are the estimated coefficients of the variables for

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common religion (that in this case is positive and statistically significant at the 1 percent level) and currency union (that losses the statistical significance). The estimated coefficient of the variable \( UPR \) also differs from that found without controls for multilateral resistance terms. With the inclusion of CFE the estimated coefficient is positive (0.200) and highly statistically significant. Results including time-varying fixed effects for exporters and importers (CYFE) reinforces this finding (column 3). When we properly account for the fact that multilateral resistance may change over time, the variable of interest presents an estimated coefficient that raises its value from 0.200 to 0.347. Thus, we find that unilateral preference regimes are associated with an increase of exports from developing countries benefiting from these nonreciprocal preference schemes.

Columns 1 to 3 report the results for three specifications that include a catch-all \( UPR \) dummy. Eicher and Henn (2009), in a recent paper on a related strand of the empirical gravity literature (the measurement of the effect of currency unions on trade), show the importance of splitting the catch-all PTA and CU dummies into the individuals PTAs and CU arrangements. According to these authors, if individual PTAs and CUs do not generate identical trade benefits, as a large empirical literature has documented, estimating an average coefficient using catch-all PTA or CU dummies generates biased results. In line with this argument, it is important to estimate the gravity equation allowing for individual UPR effects.\(^9\)

The next step of the estimation process is to run the gravity equation splitting the \( UPR \) dummy into a catch-all \( GSP \) dummy and separate dummies for AGOA, \( CBI, ATPA, EBA \) and \( ACP-EU \). Columns 4 to 6 present the results using OLS, CFE and CYFE, respectively. We focus in the latter approach since it comprehensively accounts for multilateral resistance and, therefore, it is the only fully in line with the theoretical foundations of the gravity equation. We find that \( CBI, ATPA, ACP-EU \) and \( EBA \) increase exports from the developing world, as do GSP schemes. The results for \( GSP \) are consistent with Subramanian and Wei (2007), who find that the GSP extended from the North to developing

\(^9\) Since AGOA members are also GSP beneficiaries of US, before AGOA got into force the dummy variable GSPUS takes the value of 1 for all countries under the US’s GSP scheme and after that date only for non-AGOA countries under the US’s GSP scheme. The same criteria are followed for the GSPEU dummy with respect to the dummies ACP-EU and EBA.
countries boosts trade. However, we do not find evidence that membership in the AGOA has had a positive effect on African exports to US.

Table 2 repeats the estimations with higher levels of disaggregation of the variables of interest. The first three columns present the results when we split the GSP dummy into three dummies: GSPEU (for the EU), GSPUS (for the US) and one separate dummy variable to capture all “other” GSP schemes (OtherGSP). The estimated coefficients for GSP schemes of EU and US are both positive and statistically significant, once we control for time-varying multilateral resistance terms. In particular, the estimated coefficients for GSPUS and GSPEU are 0.830 and 0.552, respectively (and they are statistically significant at the 1 percent level). However, the estimated coefficient of the variable that captures the impact of all other GSP schemes altogether is lower (0.085) and non-statistically significant at conventional levels. According to these results, the positive effect of being in the standard EU’s GSP scheme is larger than that of being in ACP-EU agreement and that of being in EBA. However, in contrast with the estimates for the impact of the US’s GSP scheme, as well as the CBI and ATPA initiatives, but in line with our previous estimates, countries belonging to AGOA do not export more to the US.

Next, we re-estimate the gravity equation including a separate dummy for each individual UPR (columns 4 to 6). It allows us to check that there is a large heterogeneity in the impact of the different GSP schemes. According to the results reported in column 6 (CYFE), the largest estimated coefficients are

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10 There are few studies that have attempted to estimate the overall effect of the GSP. Rose (2004) and Tomz, Goldstein and Rivers (2007) find a positive effect but treat the average of two-way bilateral trade as the dependent variable. Subramanian and Wei (2007) criticize Rose for averaging imports and exports, though GSP effects should differ according to whether the importer or the exporter was the recipient of the preferences. These authors, using unidirectional trade data, show an estimated effect for the GSP very similar to that reported by Rose (2004). In contrast, Goldstein, Rivers and Tomz (2007), using also a data set comprised of directed dyads, find a negative coefficient for GSP.

11 Some papers provide arguments for the ineffectiveness of AGOA. For instance, Matoo, Roy and Subramanian (2002) outline the relevance of rules of origin as a factor that could limit the achievement of high benefits. Nouve (2005) asserts that, while the AGOA apparel preferences may increase African textile and apparel exports to US, resource allocation probably would induce the reduction of overall exports. Finally, Brenton and Hoppe (2006) argue that the impact of AGOA would be enhanced if preferences were extended to all products. Additionally, these authors point out that for the majority of beneficiaries the value of preferences is very small.

12 Nilsson (2002), in a comparative analysis of the effects of the EU’s Lomé Convention and GSP on exports of developing countries, finds positive and statistically significant export effects of both, but concludes that the export impact of the Lomé Convention were greater over the period 1973-1992.
found for GSPUS (0.831) and GSPJapan (0.791), followed by GSPEU (0.559). However, our results suggest that remaining GSP schemes are not associated with an increase in exports. In fact, surprisingly, the dummy variables for Australia, New Zealand’s and Russia’s GSP schemes show a negative and statistically significant coefficient.

Column 6 of Table 2 reports the results for a specification that include catch-all PTA or CU dummies. Following Eicher and Henn (2009), we also report the results allowing for individual plurilateral PTAs and individual CUs effects (column 7). The estimated coefficients of these variables and the fixed effects are not reported in the table for ease of presentation. As we can observe, the estimated coefficients do not change in a significant way and, in particular, the estimated coefficients of the variables of interest remain nearly unaltered with two exceptions: GSPNorway and GSPSwitzerland (that are positive and now reach the statistical significance at least at the 10 percent level).

The problem of all the above estimations is that in those regressions we use the sample of countries with positive trade volumes between them. Disregarding countries that do not trade with each other may produce biased estimates (HMR, 2008). Therefore, now we turn to the analysis of the results using the two stages estimation procedure suggested by HMR (2008). Table 3 reports the results. Since our sample has time dimension we include in this framework country year fixed effects in order to capture the time-varying nature of trade costs in panel data. The results for the probit regression are presented in column 1. Before discussing the empirical results, it is worth noting that the estimation of equation (2) might be subject to the incidental parameter problem, introducing a bias in the coefficients of the rest of variables

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13 Our sample includes more than 200 individual bilateral and plurilateral PTAs and CUs. For bilateral PTAs we have estimated an average coefficient using a catch-all dummy. The inclusion of individual dummies for bilateral PTAs does not affect the results in any significant way.

14 HMR (2008) applies their two stages estimation procedure to data from 1986 including in the regression exporting and importing CFE. The working paper version of this article (HMR, 2007) also presented the results for a large sample that covered all the 1980s. However, they also used in these regressions CFE and year fixed effects instead of CYFE.

15 Following HMR (2008) we also have country pairs whose characteristics are such that their probability of trade is indistinguishable from 1. Therefore, we assign the same $\rho_{ij}$ to those country pairs with an estimated $\rho_{ij} > 0.9999999$. 
(X\textsubscript{ij} and Z\textsubscript{ij}). However, as pointed out by Fernández-Val (2007), this bias does not affect the estimated marginal effects and, therefore, the predicted values obtained for the dependent variable. These results compared with those found using CYFE in Table 2 clearly show that almost the same control variables that impact export volumes in the traditional estimation with CYFE also impact the probability that country \textit{i} exports to country \textit{j}. The exception is the dummy variable \textit{CU}, which positively affects the probability of exports but do not affect the volume of exports (in any of the previous specifications).

With regard to the dummies for nonreciprocal preference regimes, the estimated marginal effect of the variables of interest are once again positive and statistically significant for \textit{EBA}, \textit{ACP-EU} and GSP schemes of EU and US, but in contrast to CYFE, this is also the case for GSP schemes of Canada, Japan, Norway, Russia and Switzerland, suggesting that being members of these initiatives raises the probability of bilateral trade from developing countries to the cited developed countries.

Using the probit regression, as explained before, we construct two variables for correcting sample selection bias and firm heterogeneity. Both the non-linear coefficient \(\delta\) and the linear coefficient for \(\vec{\eta}\) are precisely estimated. The results for the second stage can be seen in column 2 of Table 3. The variable \textit{CU} has been excluded from the estimation for identification reasons.

The estimated coefficients are in line with those found using OLS including CYFE. At this stage, we once again find a positive and significant coefficient for \textit{EBA}, \textit{ACP-EU}, \textit{GSPUS} and \textit{GSPEU}. This is also the case for \textit{GSPCanada}, \textit{GSPJapan}, \textit{GSPNorway}, \textit{GSPSwitzerland} and \textit{GSPTurkey}, but not for AGOA, \textit{CBI}, \textit{ATPA} and the remaining GSP schemes. In particular, the largest estimated coefficients are found for \textit{GSPJapan} (0.811) and \textit{GSPUS} (0.766), which suggests that, other things equal, developing countries benefiting from trade preferences under the Japan and United States GSP schemes export more than twice to these markets. Once again, the estimated coefficients for \textit{GSPAustralia}, \textit{GSPNZ} and \textit{GSPRussia} are negative and statistically significant.

\footnote{The result for the variable \textit{CU} (currency unions) contrast with the large evidence emerged after the Rose (2000) seminal paper. See Gil, Llorca and Martínez-Serrano (2008c) for a review of the literature about the CU trade effects.}

\footnote{Following HMR (2007, footnote 26), we have also used the variable common religion for this purpose. It yields very similar results (available from the authors upon request).}
Finally, it is important to analyze whether or not unilateral preference regimes have also had an effect on exports from the preference-granting countries to the beneficiary countries. In order to study the potential effect of unilateral preference regimes on exports from developed countries to developing countries, we have augmented the gravity equation (1) by adding a binary dummy variable \( MUPR \) which is unity if \( i \) is a benefactor country of an unilateral preference regime and \( j \) is the corresponding beneficiary country. The results from a theoretically motivated specification of the gravity equation (using CYFE) appear in column 1 of Table 4. The estimated coefficients change very little with respect to those reported in column 3 of Table 1. In particular, the evidence about the positive impact of unilateral preference regimes on exports from developing countries to industrialized nations remains unaltered and the estimated coefficient of this variable is very similar to that offered in Table 1. The novelty is that we also find a positive (0.451) and statistically significant coefficient (at the 1 percent level) for the variable that captures the effect on exports from developed countries to developing countries \( (MUPR) \). It suggests that these preferential regimes have had economic success in terms of trade in both directions.

Columns 2 to 3 of Table 4 present the results for different levels of disaggregation of the UPRs dummies (using again CYFE in all the cases). Similar to the column 1 results, the coefficient estimates of the variables of interest in columns 2 to 4 of Table 4 lead to conclude that there is evidence of a positive effect on benefactor exports to developing countries. In most of the cases the unilateral programs have stimulated trade in both directions. However, and in line with our previous estimates, AGOA is an exception. It has not increased exports in any direction.

5. Conclusions

This paper investigates whether and to what extent unidirectional trade agreements have increased developing countries’ exports to richer countries. Moreover, it also analyses the potential impact in the reverse direction. Using traditional estimation techniques and recent developments in the econometric analysis of the gravity equation over the period 1990-2008, we find robust evidence that, on the whole, nonreciprocal preference regimes and GSP
schemes have had an economically significant effect on exports. However, the estimation of catch-all dummies masks heterogeneous results for the individual schemes. In particular, we find strong evidence that the ACP-EU and GSP schemes of US and EU have had a large positive effect on developing countries exports to the corresponding developed markets (US or EU). This result also applies to EBA membership and to the GSP schemes of Canada, Japan, Norway, Switzerland and Turkey once we control for time-varying multilateral resistance terms, sample selection bias and unobservable firm heterogeneity. However, we do not find evidence that membership in the AGOA scheme has had a positive effect on exports from African countries to US and the same applies for the remaining GSP schemes.

Most economists prefer two-way tariff cuts because when tariffs are reduced in a reciprocal manner rather than in an unilateral way, countries tend to produce and export on the basis of their comparative advantage (thus exporting products that they produce relatively better and importing products that others do relatively more efficiently). Despite the “lack of reciprocity”, we find that unilateral preference regimes have also had a positive effect on exports from donor countries to developing counterparts. It suggests that, in contrast with the argument raised by critics of nonreciprocal programs, unilateral preference regimes have not had a perverse effect on trade policies of beneficiary countries (in particular with respect to the preference-granting countries).
References


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Notes: Regressand: log of real bilateral exports. Robust standard errors (clustered by country-pairs) are in parentheses. * significant at 10%; ** significant at 5%; *** significant at 1%.

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Notes: Regressand: log of real bilateral exports. Robust standard errors (clustered by country-pairs) are in parentheses. ** significant at 10%, *** significant at 5%, **** significant at 1%. The regression reported in column 7 includes individual dummies for all plurilateral PTAs and CUs.
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<td>0.392 (0.231)***</td>
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<td>0.120 (0.054)**</td>
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<td>0.214 (0.013)**</td>
<td>0.362 (0.191)**</td>
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<td>ETA</td>
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Notes: Robust standard errors (clustered by country-pairs) are in parentheses. * significant at 10%; ** significant at 5%; *** significant at 1%.

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<td>-1.219 (0.020)***</td>
<td>-1.225 (0.020)***</td>
<td>-1.234 (0.020)***</td>
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<td>0.443 (0.071)</td>
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<td>MACP-EU_{ij}</td>
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<td>0.464 (0.169)**</td>
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<td>0.848 (0.156)</td>
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Notes: Regressand: log of real bilateral exports. Robust standard errors (clustered by country-pairs) are in parentheses.* significant at 10%; ** significant at 5%; *** significant at 1%.
### Table A1: Sample of countries.

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<td>El Salvador</td>
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